## Binomial tests Chi square tests P-values Confidence Intervals Null Hypothesis testing



Biostatistics Course 2023 Lecture 2 Tuesday, 25 July 2023 1:00pm - 3:00pm **Example 1: Human Sex Ratio** 

### Computing sex ratio of humans is one of the oldest applications of statistics

_			
	year	male	female
	1629	5218	4683
	1630	4858	4457
	1631	4422	4102
	1632	4994	4590
	1633	5158	4839
	1634	5035	4820
	1635	5106	4928
	1636	4917	4605
	1637	4703	4457
-			

Arbuthnott J (1711). An Argument for Divine Providence, taken from the Constant Regularity observed in the Births of both Sexes.

https://www.openintro.org/stat/data/?data=arbuthnot

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This analysis expects that each value in the data table is an actual number of events or items, and is not normalized in any way.

#### Data set to analyze

A: year 1634

#### Enter expected values as

Actual numbers of objects or events

Percentages

With vo rows, perform

nomial test (recommended)

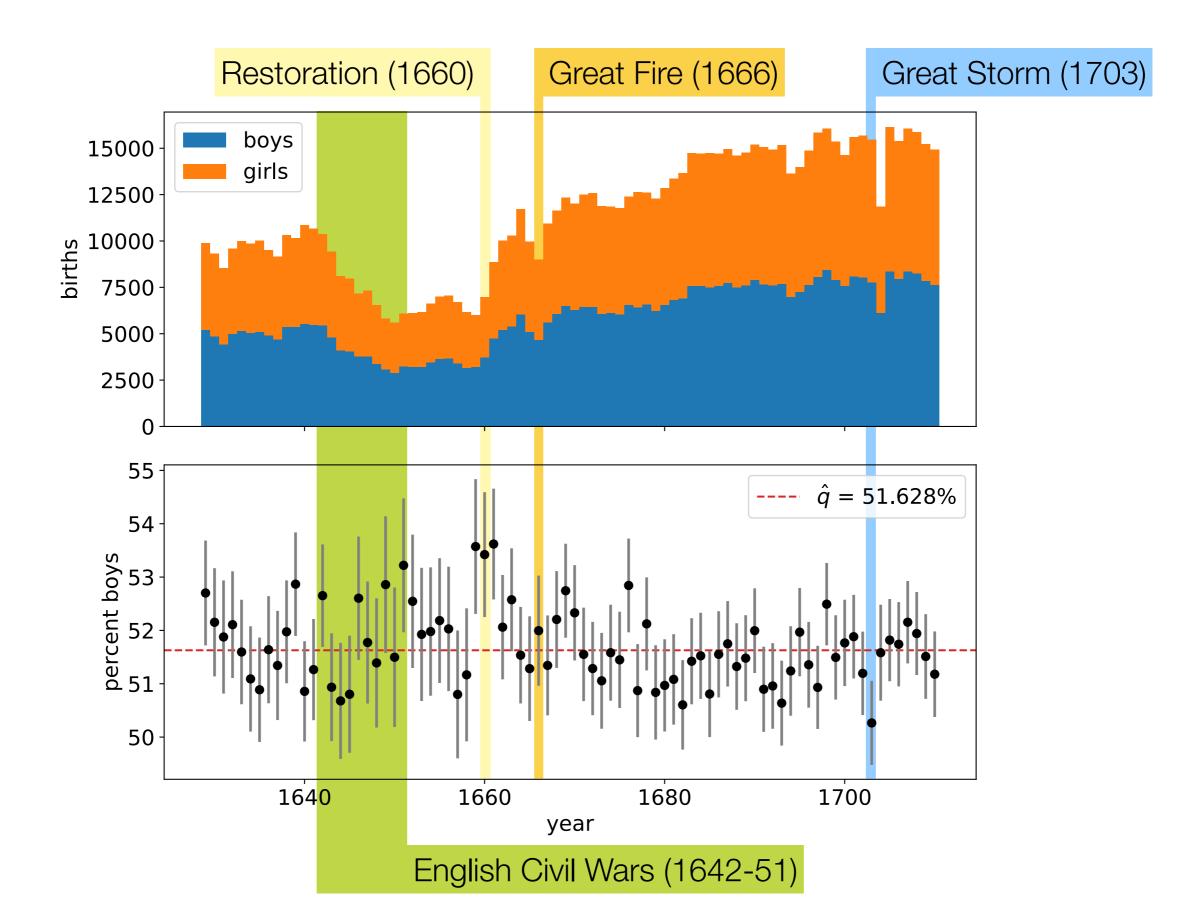
Ch-square test for goodness of fit

#### Expected distribution

	Row	Outcome	Observed %	Expected %
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	2	girls	48.91	50
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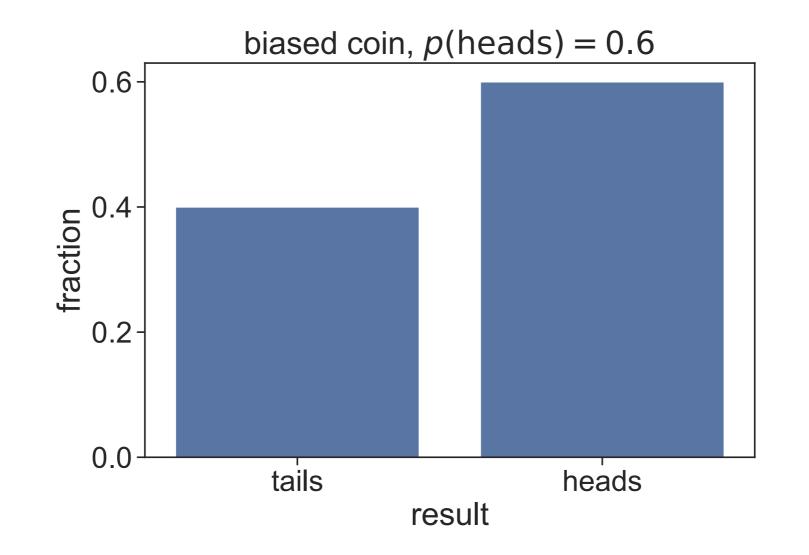
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▼ Layouts >>>	9	is discrepancy significant (F < 0.05)?	165						+
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		boys							+
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### Births in London, 1629-1710



Example 2: A biased coin

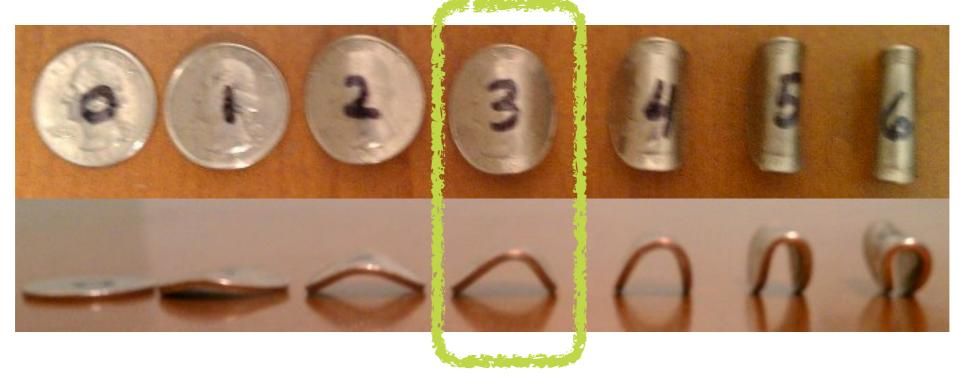
## Biased coins are modeled using a <u>Bernoulli distribution</u>, which describes probabilities for a <u>binary variable</u>



### Making a biased coin

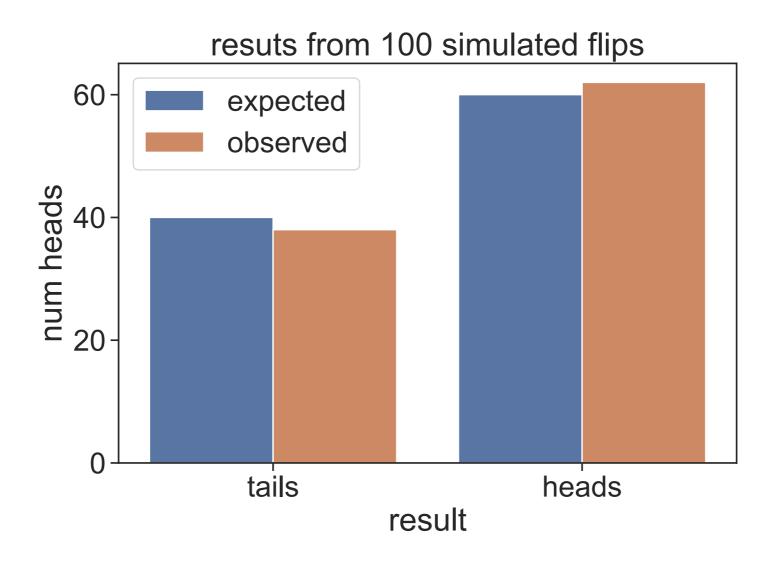


## $p(\text{heads}) \approx 60\%$



Mike Izbicki (Claremont McKenna College) <u>https://izbicki.me/blog/how-to-create-an-unfair-coin-and-prove-it-with-math.html</u>

## The number of heads after 100 flips of the biased coin will resemble the underlying probabilities, but will not match exactly



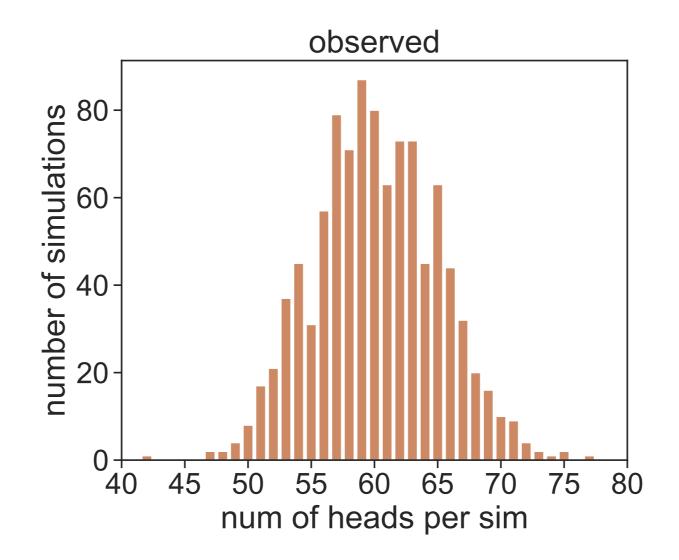
expected: 60 heads, 40 tails

observed: 62 heads, 38 tails

How much deviation from the expected values do we expect?

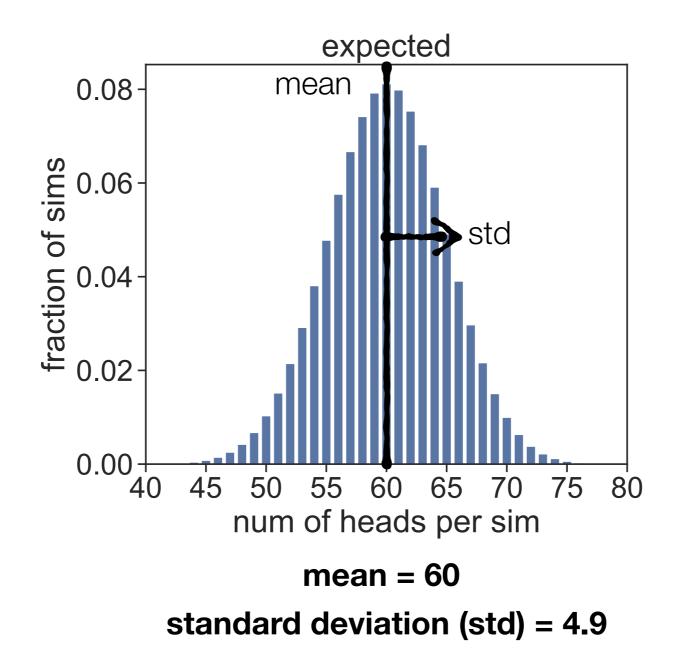
### There is substantial variation across replicates. This is to be expected.

### Results from 1000 simulations, 100 flips per simulation



## The variation in the number of heads from replicate to replicate is described by a <u>binomial distribution</u>

### Results from 1000 simulations, 100 flips per simulation



Can we determine whether or not a coin is biased by flipping it 100 times?

Suppose we flip a coin 100 times and observe 62 heads (and 38 tails).

**Null hypothesis**: heads and tails are equally likely, i.e. p(heads) = 50%

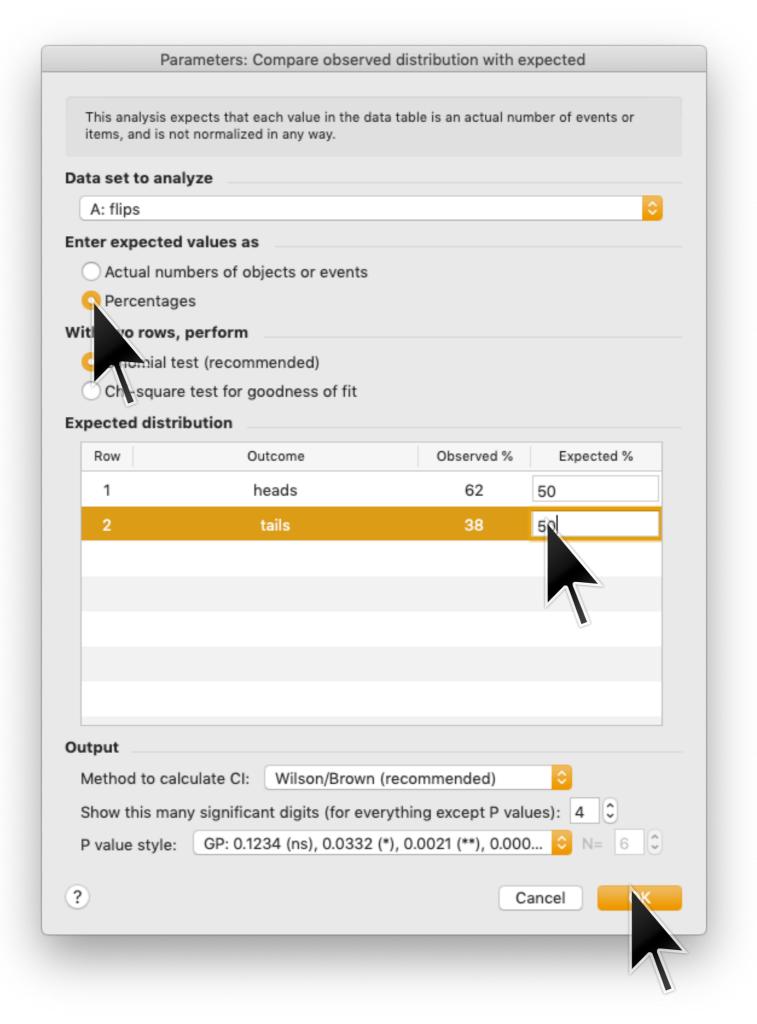
Alternative hypothesis: heads and tails are not equally likely, i.e.  $p(\text{heads}) \neq 50\%$ 

Our observation (62 heads) may or may not allow us to reject the null hypothesis and thus accept the alternative hypothesis.

No amount of data, however, can cause us to accept the null hypothesis.

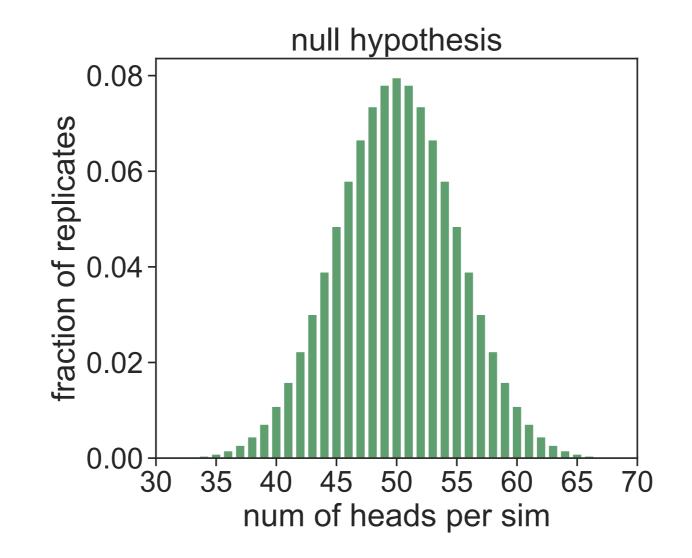
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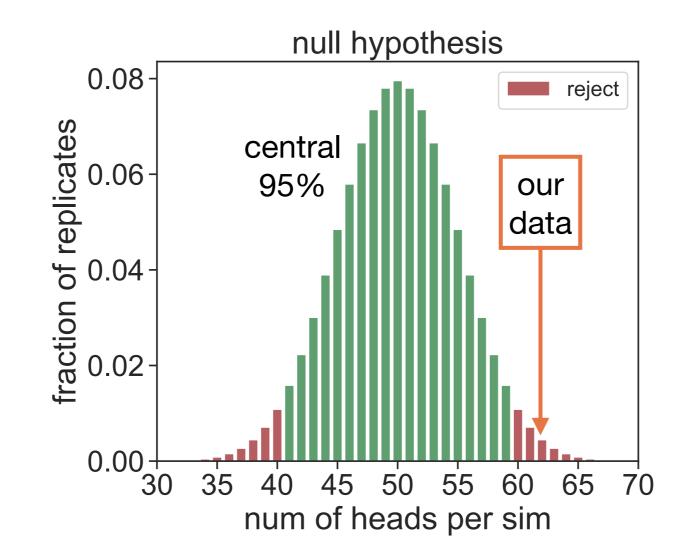
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	11	heads	50.00	62	50.00	62.00	52.21 to 70.90	
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### The null hypothesis is assessed by where the data fall within the null distribution



# We reject the null hypothesis of the data fall too far away from the bulk of the distribution

If the null hypothesis is true, data should fall within the green region 95% of the time, and within the red "reject" region 5% of the time.

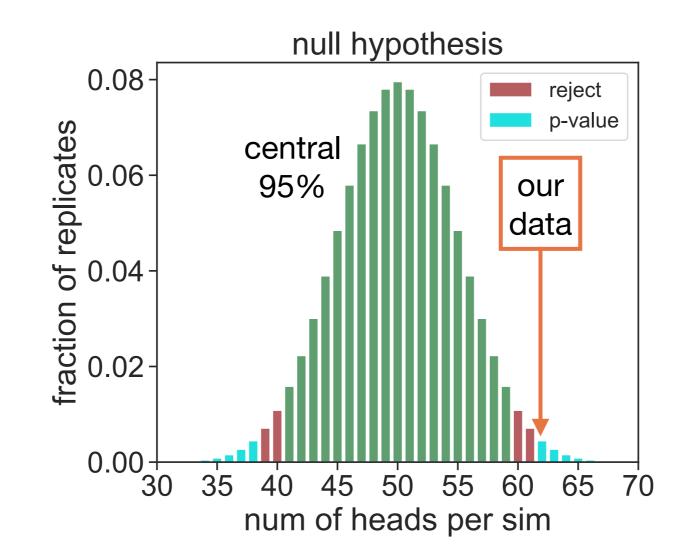


Our assumed dataset (62 heads) lies outside the central 95%.

We can therefore reject the null hypothesis with 95% confidence.

## P-values quantify the probability of data being as or more extreme than the data in hand were the null hypothesis true.

The P-value threshold of 0.05 comes from adopting a confidence threshold of 95%.

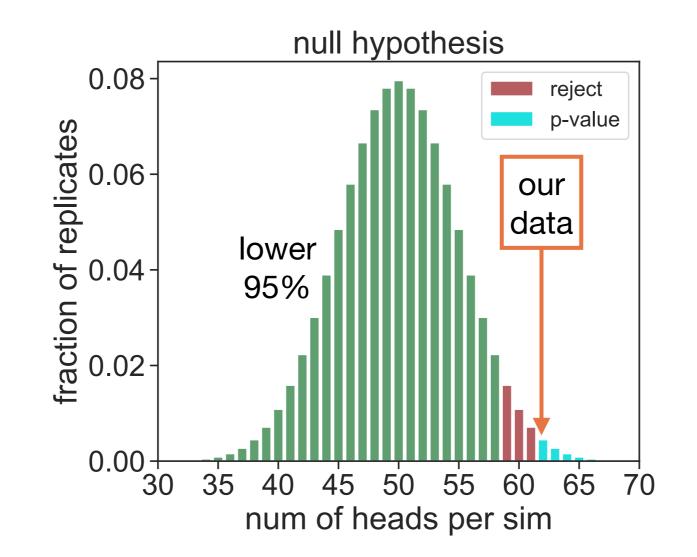


We find that **p=0.0210** for the <u>two-sided</u> test.

We therefore say that our result is "statistically significant"

# P-values quantify the probability of data being as or more extreme than the data in hand were the null hypothesis true.

A <u>one-sided hypothesis test</u> only considers one side of the distribution.

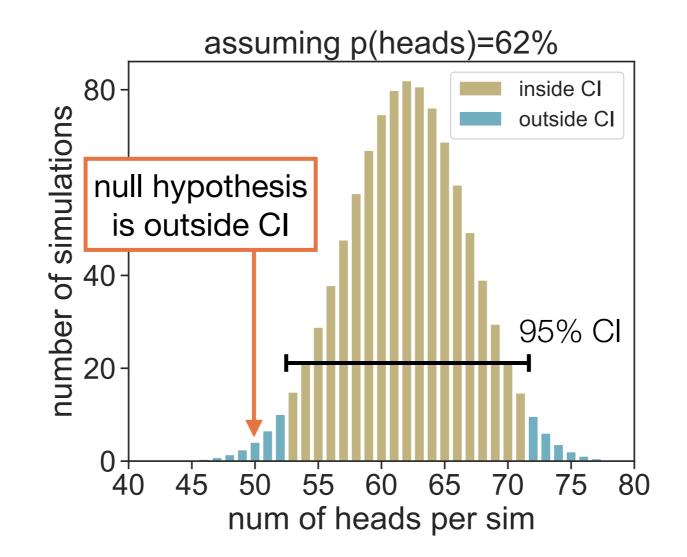


We find that **p=0.0105** for the <u>one-sided</u> test.

In general, two-sided tests are more conservative than one-sided tests.

Unless you have good reason to do otherwise, use two-sided tests.

### **Confidence intervals (CIs) are more informative than P-values**



We conclude that p(heads) lies within [52.5%, 71.5%] with 95% confidence.

We can reject the null hypothesis because it lies outside this confidence interval.

- "Statistically significant" does not actually mean "significant" in the normal sense. At best, it means "detectable".
- P-values do not say how big an observed effect is.
- P-values do not say how important that observed effect is.
- P-values calculations rely on assumptions, and violation of any of those assumptions can render P-values meaningless.
- Perhaps most severe is the fact that <u>P-values do not actually quantify you how</u> <u>likely or unlikely your null hypothesis is</u>!

- Like a P-value, a CI communicates statistical significance (i.e. detectability).
- A CI also communicates effect size, as well as the uncertainty in that effect size.
- A 95% CI does not actually mean that the true value of a parameter lies within that interval with 95% probability. Still, this (extremely common) misinterpretation is largely benign compared to the misinterpretation of P-values.
- However, P-values are more commonly reported than confidence intervals.

The perils of null hypothesis testing

Step 1: Specify a null hypothesis.

**Step 2:** Specify a confidence level (usually 95%)

Step 3: Identify the appropriate statistical test



**Result:** P-value summarizing how unlikely the data is compared to null hypothesis expectations.

Roughly speaking, P-values quantify how likely our data would be if the null hypothesis were true.

p(data | null hypothesis)

P-values <u>do not</u> quantify the probability of the null hypothesis given our data. Unfortunately, this is the quantity that we actually care about.

*p*(null hypothesis | data)

My opinion: the use of P-values to reject hypotheses is predicated on the base rate fallacy

By convention P < 0.05, then one rejects null hypothesis, supposedly because p(null hypothesis | data) is small.

For this to make sense, one has to accept the base rate fallacy, i.e.,

 $p(\text{data} | \text{null hypothesis}) \approx p(\text{null hypothesis} | \text{data})$ 

Whether or not this is true in a specific case depends on the prior odds,

*p*(null hypothesis),

which Frequentist statistics refuses to consider.

Frequentist statistics (a.k.a. classical statistics) focuses on <u>likelihood</u>:

p(data | hypothesis).

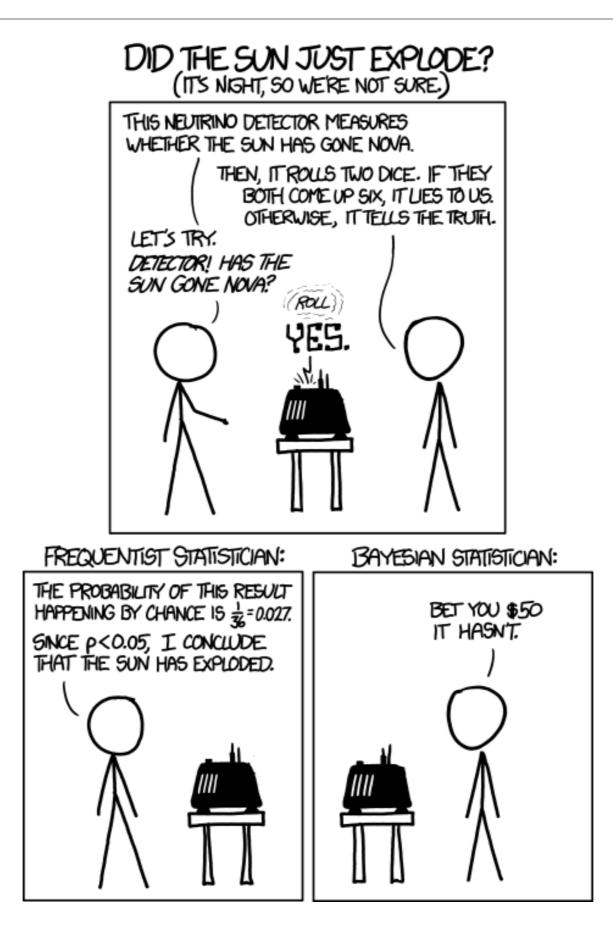
### Iron Law of Frequentist Statistics:

Never compute the probability of a hypothesis.

Bayesian statistics focuses on computing posterior probabilities:

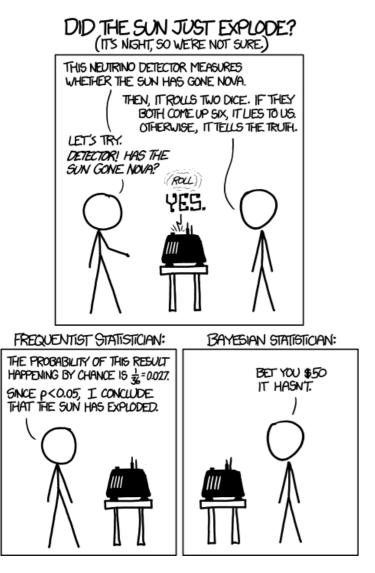
*p*(hypothesis | data).

**Example 3: Supernova detection machine** 



https://xkcd.com/1132/

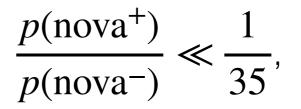
Bayes's theorem (from yesterday):



https://xkcd.com/1132/

$$\frac{p(\text{nova}^+ | \text{detector}^+)}{p(\text{nova}^- | \text{detector}^+)} = \frac{p(\text{detector}^+ | \text{nova}^+)}{p(\text{detector}^+ | \text{nova}^-)} \times \frac{p(\text{nova}^+)}{p(\text{nova}^-)}$$
$$\left[\frac{35/36}{1/36} = 35\right]$$

If our prior belief is that a supernova is very unlikely, i.e.

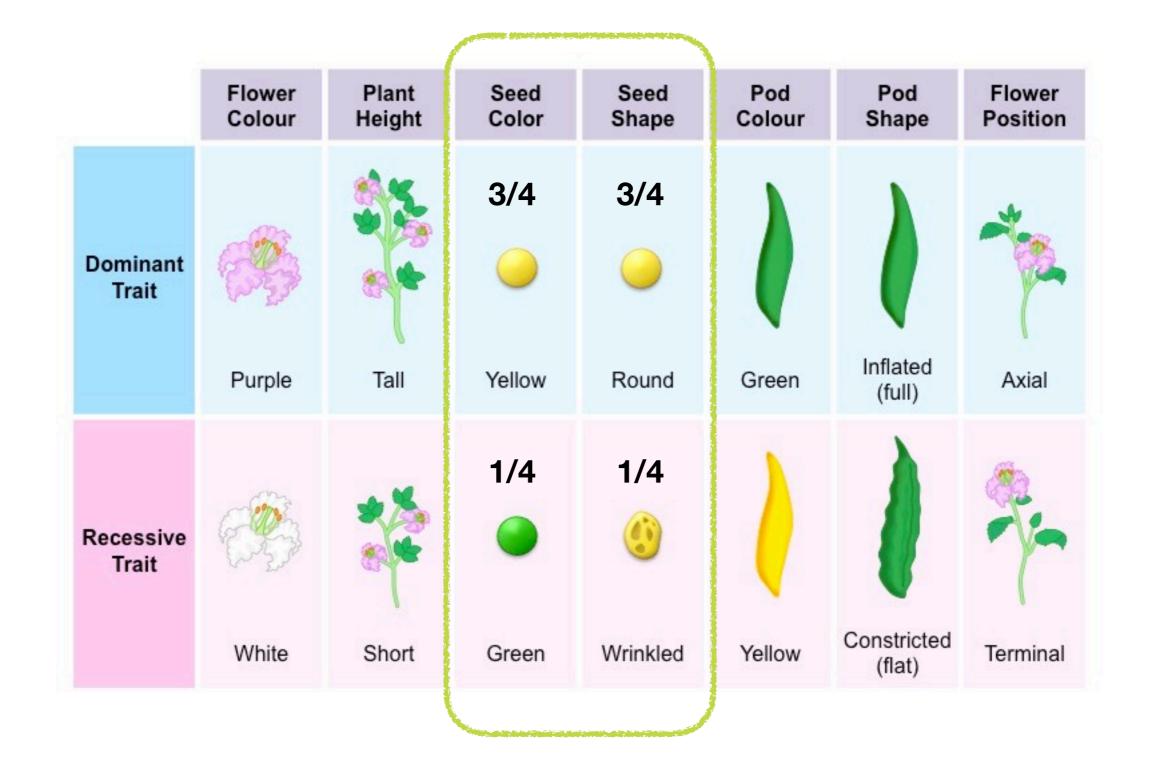


then we still shouldn't believe the sun has gone nova.

Even though, with a null hypothesis of nova-,

P value = 
$$p(detector^+ | nova^-) = \frac{1}{36} = 0.028 < 0.05$$

**Example 4: Mendel's Peas** 



https://ib.bioninja.com.au/standard-level/topic-3-genetics/34-inheritance/mendels-laws.html

### Chi square test (known proportions)

### **Example: Mendel's peas**

	observed	expected proportion	expected counts
Round & yellow	315	9/16	312.75
Round & green	108	3/16	104.25
Angular & yellow	101	3/16	104.25
Angular & green	32	1/16	34.75
Total	556	16/16	556.00

### **Null Hypothesis:**

observations in K = 4 different categories occur in the expected proportions

Data: number of observations in each category

Statistic: 
$$\chi^2 = \sum \frac{(\text{observed} - \text{expected})^2}{\text{expected}}$$

**Null distribution:** Chi square distribution with K - 1 = 3 degrees of freedom (DOF)

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Parameters: Compare observed distribution with expected

This analysis expects that each value in the data table is an actual number of events or items, and is not normalized in any way.

#### Data set to analyze

A: observed

#### Enter expected values as

Actual numbers of objects or events

**Q**Percentages

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Ch-square test for goodness of fit

#### **Expected distribution**

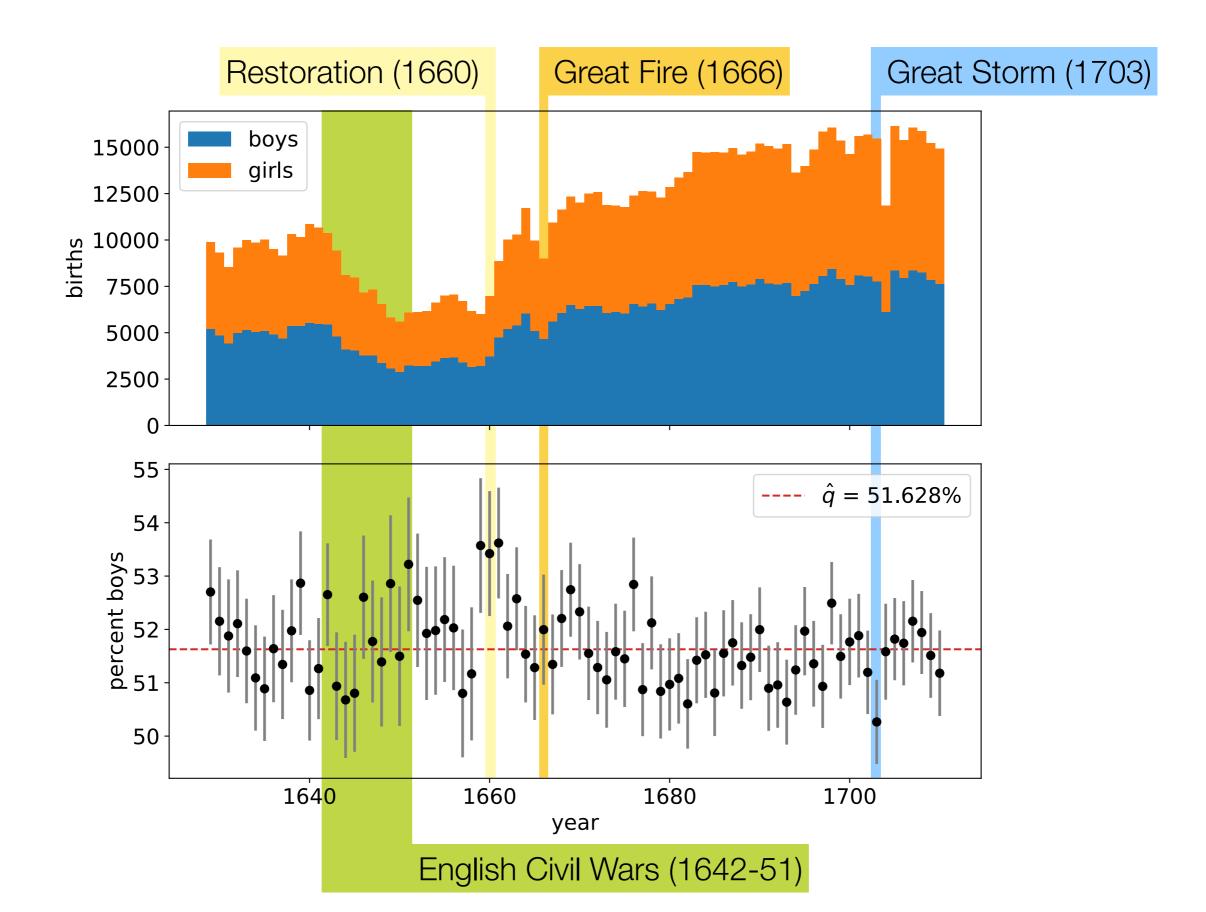
Row	Outcome	Observed %	Expected %
1	Round & yellow!	56.65	56.25
2	Round & green	19.42	18.75
3	Angular & yellow!	18.17	18.75
4	Angular & green!	5.76	6.25
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	1:	Round & green	104.3	108	18.75	19.42
	14	Angular & yellow	104.3	101	18.75	18.17
	1	Angular & green	34.75	32	6.250	5.755
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**Example 4: Human sex ratio in London over time** 

### Is it possible that the boy/girl ratio changes from year to year?



### Chi square test (unknown proportions)

	male	female
1629	5218	4683
1630	4858	4457
1631	4422	4102
1632	4994	4590
1633	5158	4839
1634	5035	4820
1635	5106	4928
1636	4917	4605
1637	4703	4457
1638	5359	4952

### sex

### **Null Hypothesis:**

Two multi-category variables A and B are independent, i.e.,  $p(A, B) = p(A) \cdot p(B)$ 

### **Statistic:**

$$\chi^2 = \sum \frac{(\text{observed} - \text{expected})^2}{\text{expected}}$$

Null distribution:Chi square distribution withDOF = nm - m - n + 1wherem = number of possible values for An = number of possible values or B

year

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Project info 1		2 1630	4858	4457						
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	-	9 1637	4703	4457						
	-	10 1638	5359	4952						
	-	11 1639	5366	4784						
	-	12 1640	5518	5332						
	-	13 1641	5470	5200						
		14 1642	5460	4910						
Family	>>	15 1643	4793	4617						
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		17 1645	4047	3919						
		18 1646	3768	3395						
		19 1647	3796	3536						
		20 1648	3363	3181						
		21 1649	3079	2746						
		22 1650	2890	2722						
	-	23 1651	3231	2840						
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Transform	B:girls
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Used fo	prospective a	nd experimenta	al studies		
🗌 Differen	ce between pro	oportions (attri	outable risk)	and NNT	
Used fo	prospective a	nd experimenta	al studies		
🗌 Odds ra	tio				
Used fo	retrospective	case-control s	tudies		
🗌 Sensitiv	ty, specificity a	and predictive v	alues		
Used fo	diagnostic tes	sts			
Method to c	ompute the P	value			
O Fisher's	exact test				
🔿 Yates' d	ontinuity corre	cted chi-squar	e test		
🜔 Chi-squ	are test				
hi-squ	are test for tre	nd			
		pare proportions ion). The chi-squ		chi-square test (w s are equivalent.	vith
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arbuthnot					
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▼ Info ③ Project info 1	2				
⊕ New Info	3	P value and statistical significance			
	> 4	Test	Chi-square		
Contingency of arbuthnot	5	Chi-square, df	169.7, 81		
① New Analysis	6	P value	<0.0001		
	» 0 7		<0.0001 ****		
<ul> <li>⊕ New Graph</li> <li>▼ Layouts</li> </ul>		P value summary			
New Layout	0	One- or two-sided	NA		
	9	Statistically significant (P < 0.05)?	Yes		
	10				
	11	Data analyzed			
	12	Number of rows	82		
	13	Number of columns	2		
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