Binomial tests Chi square tests P-values Confidence Intervals Null Hypothesis testing



Biostatistics Course 2024 Lecture 2 Tuesday, 9 July 2024 10:00am - 12:00pm **Example 1: Human Sex Ratio**

Computing sex ratio of humans is one of the oldest applications of statistics

year	male	female
1629	5218	4683
1630	4858	4457
1631	4422	4102
1632	4994	4590
1633	5158	4839
1634	5035	4820
1635	5106	4928
1636	4917	4605
1637	4703	4457
	1629 1630 1631 1632 1633 1634 1635 1636	1629521816304858163144221632499416335158163450351635510616364917

Arbuthnott J (1711). An Argument for Divine Providence, taken from the Constant Regularity observed in the Births of both Sexes.

https://www.openintro.org/stat/data/?data=arbuthnot

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This analysis expects that each value in the data table is an actual number of events or items, and is not normalized in any way.

Data set to analyze

A: year 1634

Enter expected values as

Actual numbers of objects or events

Percentages

With vo rows, perform

nomial test (recommended)

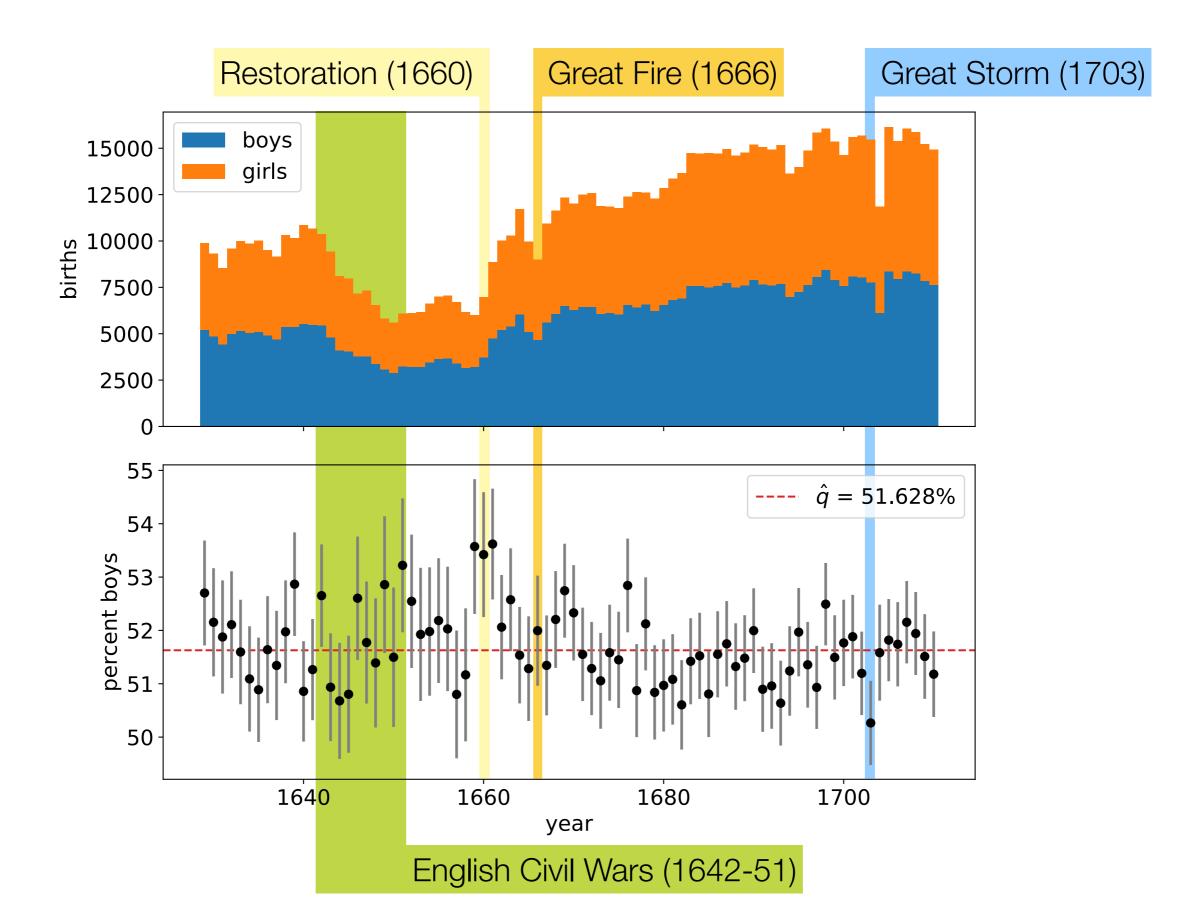
Ch-square test for goodness of fit

Expected distribution

	Row	Outcome	Observed %	Expected %
	1	boys	51.09	50
	2	girls	48.91	50
Out	tput			
Ν	lethod to calc	ulate CI: Wilson/Brow	n (recommended)	0
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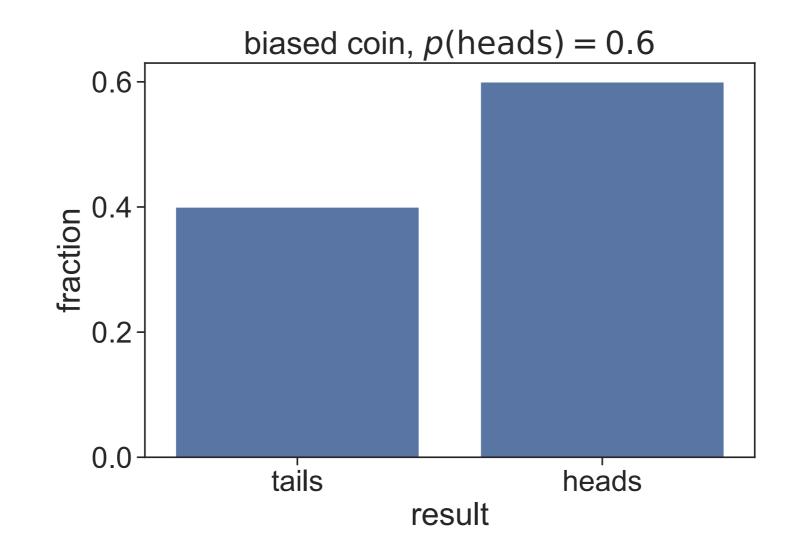
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	11		4928	5035	50.00	51.09	50.10 to 52.08	veu /o	+
		boys							+
	12	girls	4928	4820	50.00	48.91	47.92 to 49.90		 +
	13	TOTAL	9855	9855	100.0	100.00			 +
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Births in London, 1629-1710



Example 2: A biased coin

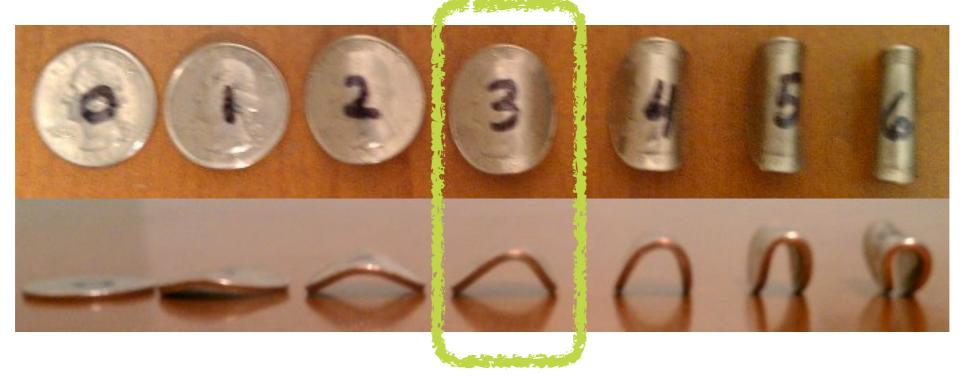
Biased coins are modeled using a <u>Bernoulli distribution</u>, which describes probabilities for a <u>binary variable</u>



Making a biased coin

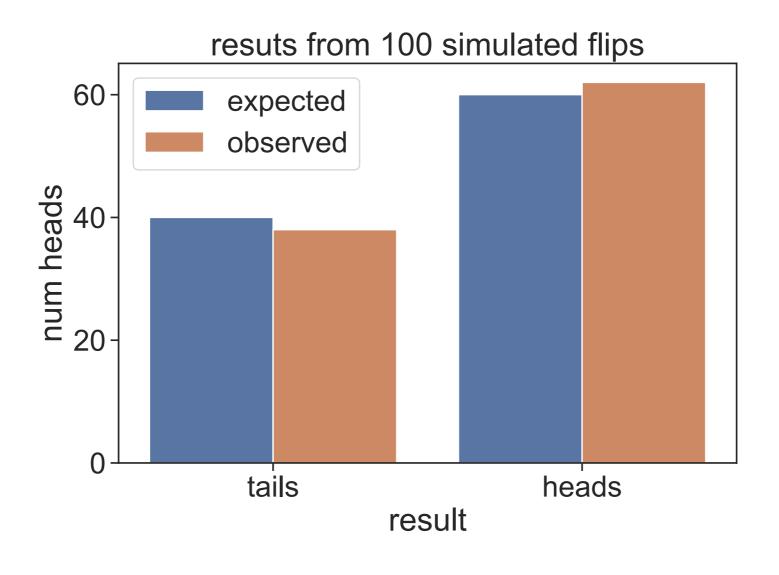


$p(\text{heads}) \approx 60\%$



Mike Izbicki (Claremont McKenna College) <u>https://izbicki.me/blog/how-to-create-an-unfair-coin-and-prove-it-with-math.html</u>

The number of heads after 100 flips of the biased coin will resemble the underlying probabilities, but will not match exactly



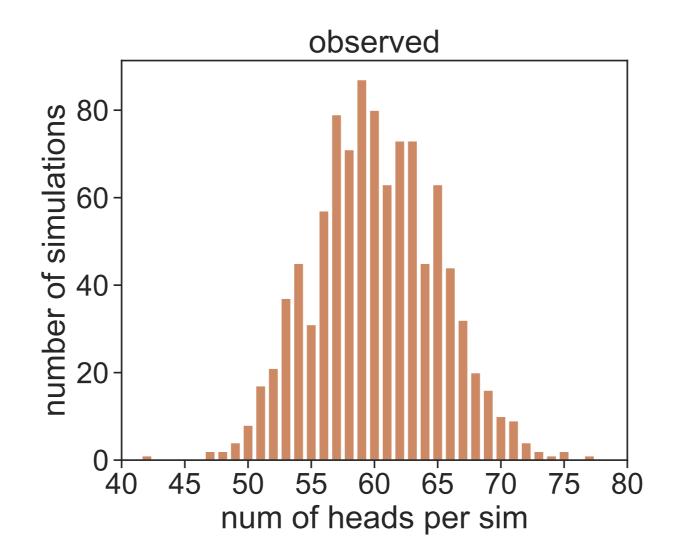
expected: 60 heads, 40 tails

observed: 62 heads, 38 tails

How much deviation from the expected values do we expect?

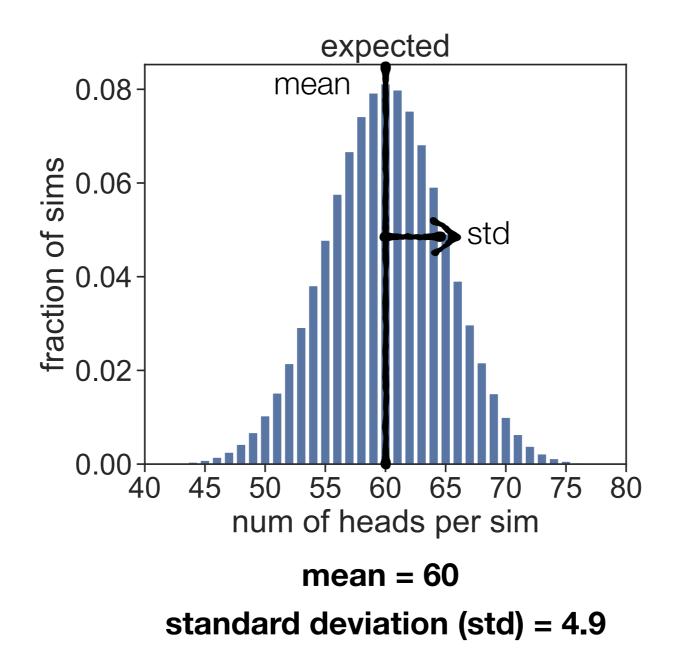
There is substantial variation across replicates. This is to be expected.

Results from 1000 simulations, 100 flips per simulation



The variation in the number of heads from replicate to replicate is described by a <u>binomial distribution</u>

Results from 1000 simulations, 100 flips per simulation



Can we determine whether or not a coin is biased by flipping it 100 times?

Suppose we flip a coin 100 times and observe 62 heads (and 38 tails).

Null hypothesis: heads and tails are equally likely, i.e. p(heads) = 50%

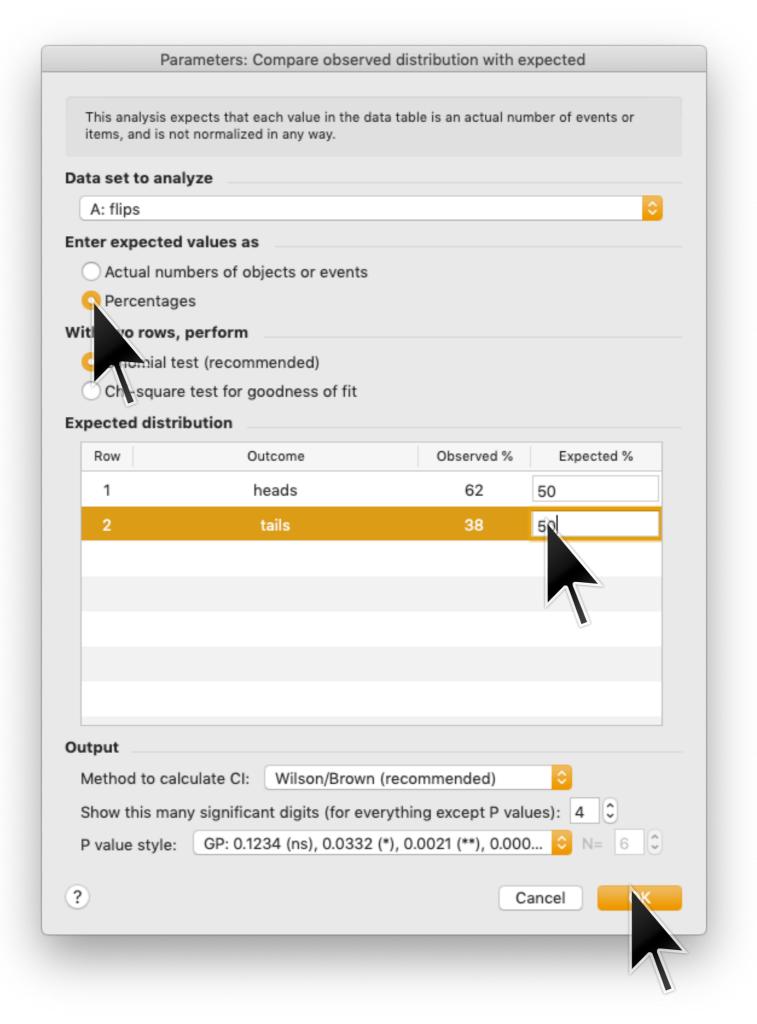
Alternative hypothesis: heads and tails are not equally likely, i.e. $p(\text{heads}) \neq 50\%$

Our observation (62 heads) may or may not allow us to reject the null hypothesis and thus accept the alternative hypothesis.

No amount of data, however, can cause us to accept the null hypothesis.

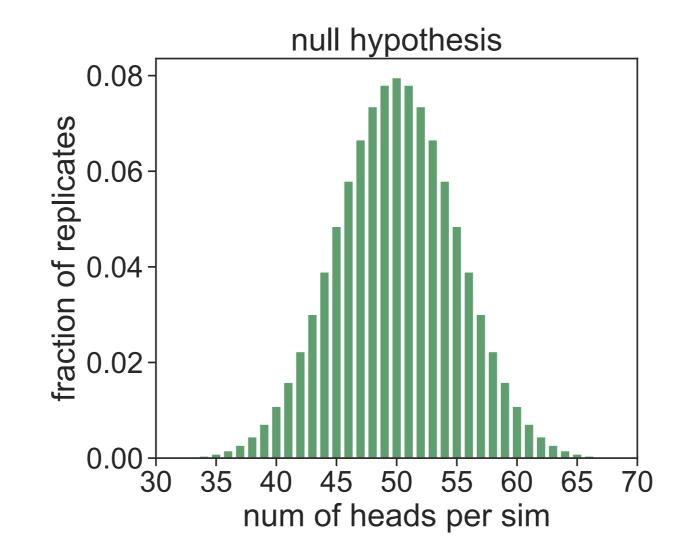
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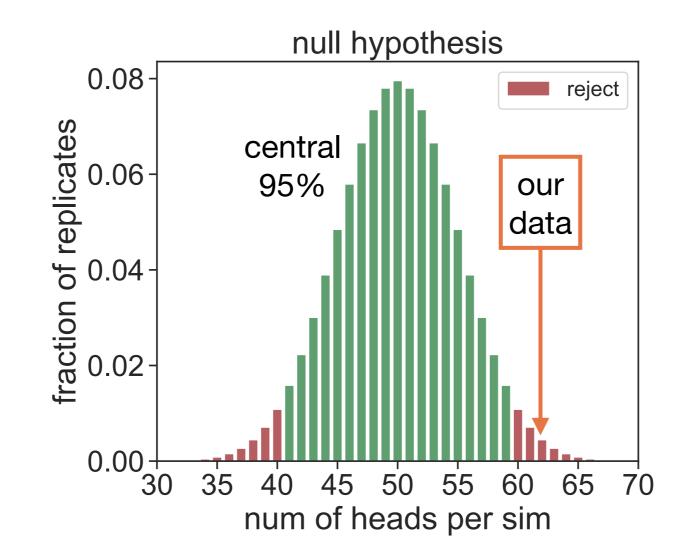
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▼ Graphs >>>	6	P (two-tailed)	0.0210					
🗠 Data 1	7	P value summary	*					
+ New Graph	8	Is discrepancy significant (P < 0.05)?	Yes					
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① New Layout	10	Outcome	Expected #	Observed #	Expected %	Observed %	95% CI of Observed %	
	11	heads	50.00	62	50.00	62.00	52.21 to 70.90	
0	12	tails	50.00	38	50.00	38.00	29.10 to 47.79	
Family >>	13	TOTAL	100.0	100.0	100.0	100.00		
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The null hypothesis is assessed by where the data fall within the null distribution



We reject the null hypothesis of the data fall too far away from the bulk of the distribution

If the null hypothesis is true, data should fall within the green region 95% of the time, and within the red "reject" region 5% of the time.

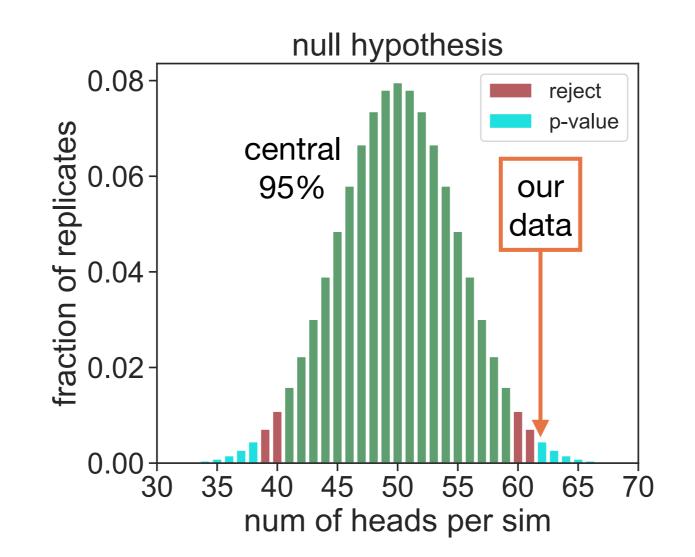


Our assumed dataset (62 heads) lies outside the central 95%.

We can therefore reject the null hypothesis with 95% confidence.

P-values quantify the probability of data being as or more extreme than the data in hand were the null hypothesis true.

The P-value threshold of 0.05 comes from adopting a confidence threshold of 95%.

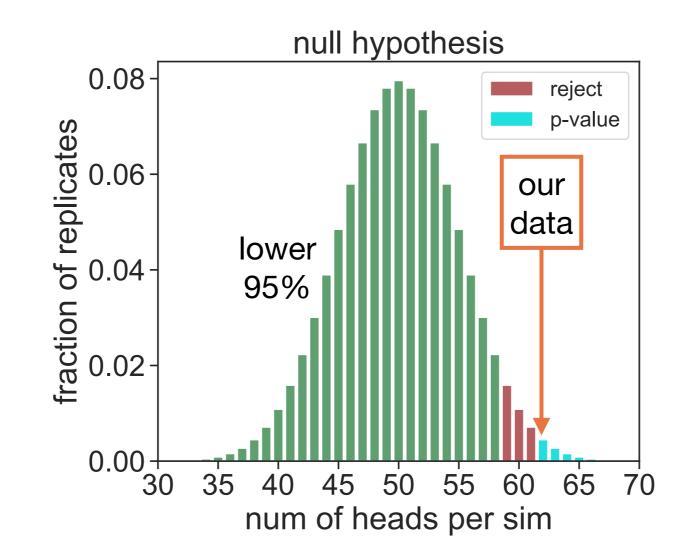


We find that **p=0.0210** for the <u>two-sided</u> test.

We therefore say that our result is "statistically significant"

P-values quantify the probability of data being as or more extreme than the data in hand were the null hypothesis true.

A <u>one-sided hypothesis test</u> only considers one side of the distribution.

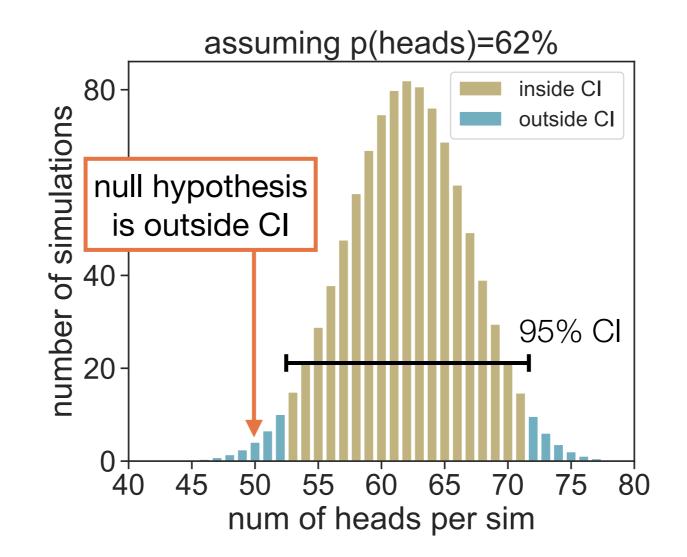


We find that **p=0.0105** for the <u>one-sided</u> test.

In general, two-sided tests are more conservative than one-sided tests.

Unless you have good reason to do otherwise, use two-sided tests.

Confidence intervals (CIs) are more informative than P-values



We conclude that p(heads) lies within [52.5%, 71.5%] with 95% confidence.

We can reject the null hypothesis because it lies outside this confidence interval.

- "Statistically significant" does not actually mean "significant" in the normal sense. At best, it means "detectable".
- P-values do not say how big an observed effect is.
- P-values do not say how important that observed effect is.
- P-values calculations rely on assumptions, and violation of any of those assumptions can render P-values meaningless.
- Perhaps most severe is the fact that <u>P-values do not actually quantify you how</u> <u>likely or unlikely your null hypothesis is</u>!

- Like a P-value, a CI communicates statistical significance (i.e. detectability).
- A CI also communicates effect size, as well as the uncertainty in that effect size.
- A 95% CI does not actually mean that the true value of a parameter lies within that interval with 95% probability. Still, this (extremely common) misinterpretation is largely benign compared to the misinterpretation of P-values.
- However, P-values are more commonly reported than confidence intervals.

The perils of null hypothesis testing

Step 1: Specify a null hypothesis.

Step 2: Specify a confidence level (usually 95%)

Step 3: Identify the appropriate statistical test



Result: P-value summarizing how unlikely the data is compared to null hypothesis expectations.

Roughly speaking, P-values quantify how likely our data would be if the null hypothesis were true.

p(data | null hypothesis)

P-values <u>do not</u> quantify the probability of the null hypothesis given our data. Unfortunately, this is the quantity that we actually care about.

p(null hypothesis | data)

My opinion: the use of P-values to reject hypotheses is predicated on the base rate fallacy

By convention P < 0.05, then one rejects null hypothesis, supposedly because p(null hypothesis | data) is small.

For this to make sense, one has to accept the base rate fallacy, i.e.,

 $p(\text{data} | \text{null hypothesis}) \approx p(\text{null hypothesis} | \text{data})$

Whether or not this is true in a specific case depends on the prior odds,

p(null hypothesis),

which Frequentist statistics refuses to consider.

Frequentist statistics (a.k.a. classical statistics) focuses on <u>likelihood</u>:

p(data | hypothesis).

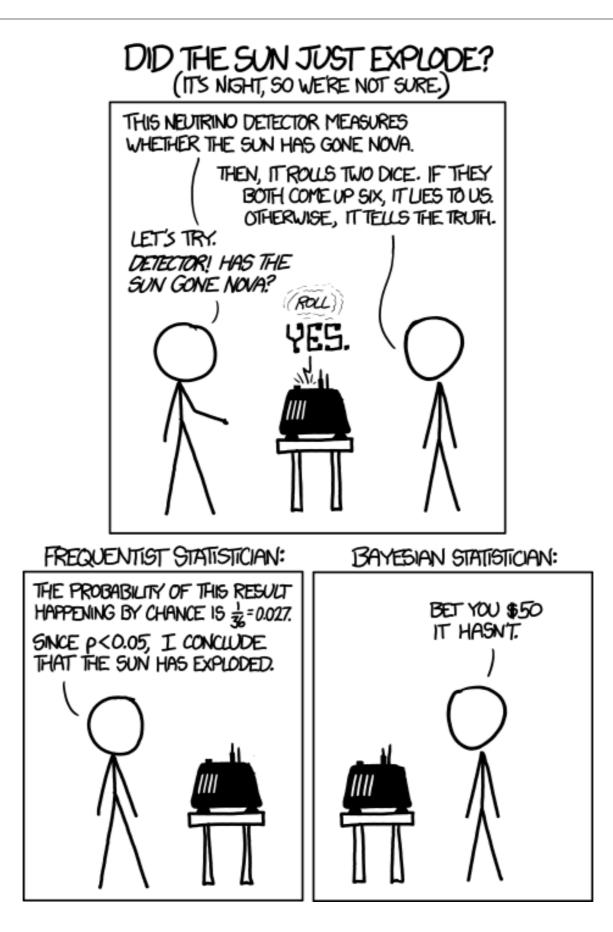
Iron Law of Frequentist Statistics:

Never compute the probability of a hypothesis.

Bayesian statistics focuses on computing posterior probabilities:

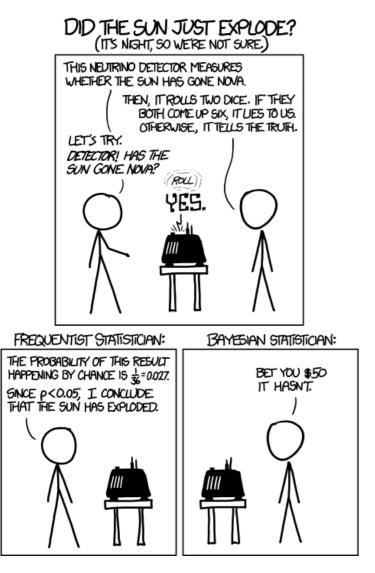
p(hypothesis | data).

Example 3: Supernova detection machine



https://xkcd.com/1132/

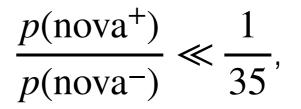
Bayes's theorem (from yesterday):



https://xkcd.com/1132/

$$\frac{p(\text{nova}^+ | \text{detector}^+)}{p(\text{nova}^- | \text{detector}^+)} = \frac{p(\text{detector}^+ | \text{nova}^+)}{p(\text{detector}^+ | \text{nova}^-)} \times \frac{p(\text{nova}^+)}{p(\text{nova}^-)}$$
$$\left[\frac{35/36}{1/36} = 35\right]$$

If our prior belief is that a supernova is very unlikely, i.e.

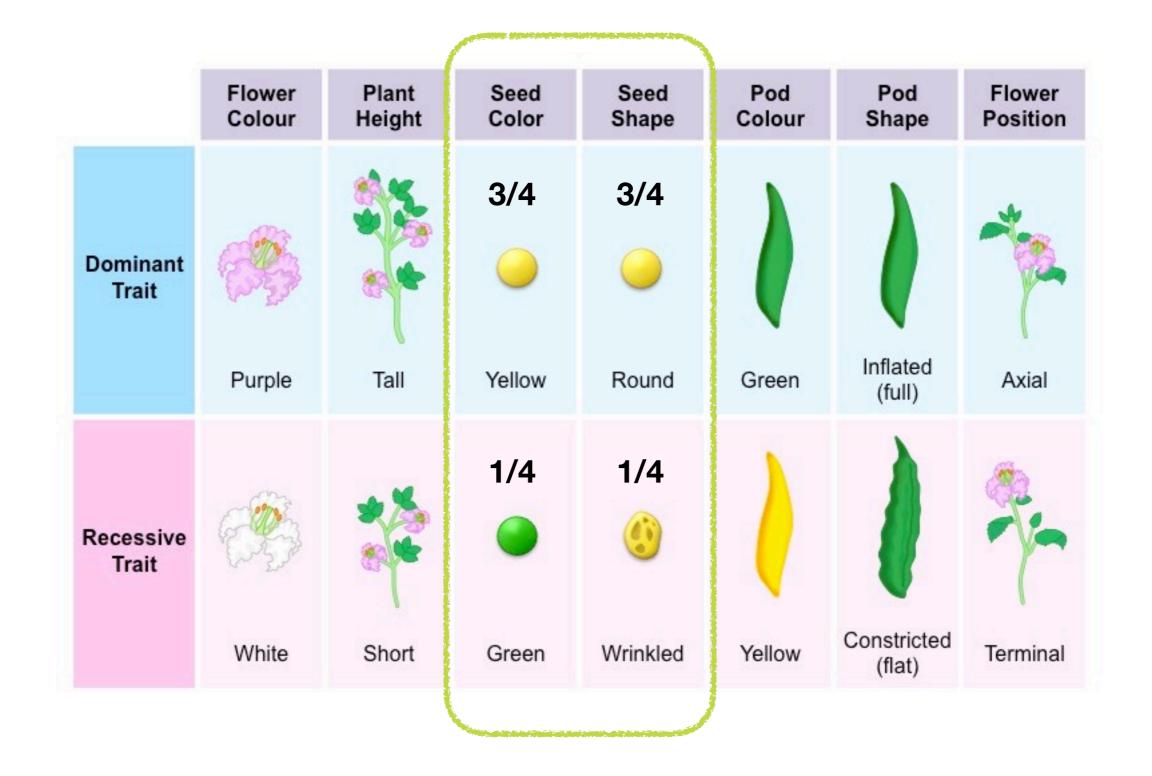


then we still shouldn't believe the sun has gone nova.

Even though, with a null hypothesis of nova-,

P value =
$$p(detector^+ | nova^-) = \frac{1}{36} = 0.028 < 0.05$$

Example 4: Mendel's Peas



https://ib.bioninja.com.au/standard-level/topic-3-genetics/34-inheritance/mendels-laws.html

Chi square test (known proportions)

Example: Mendel's peas

	observed	expected proportion	expected counts
Round & yellow	315	9/16	312.75
Round & green	108	3/16	104.25
Angular & yellow	101	3/16	104.25
Angular & green	32	1/16	34.75
Total	556	16/16	556.00

Null Hypothesis:

observations in K = 4 different categories occur in the expected proportions

Data: number of observations in each category

Statistic:
$$\chi^2 = \sum \frac{(\text{observed} - \text{expected})^2}{\text{expected}}$$

Null distribution: Chi square distribution with K - 1 = 3 degrees of freedom (DOF)

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Parameters: Compare observed distribution with expected

This analysis expects that each value in the data table is an actual number of events or items, and is not normalized in any way.

Data set to analyze

A: observed

Enter expected values as

Actual numbers of objects or events

QPercentages

With yo rows, perform

a romaial test (recommended)

Ch-square test for goodness of fit

Expected distribution

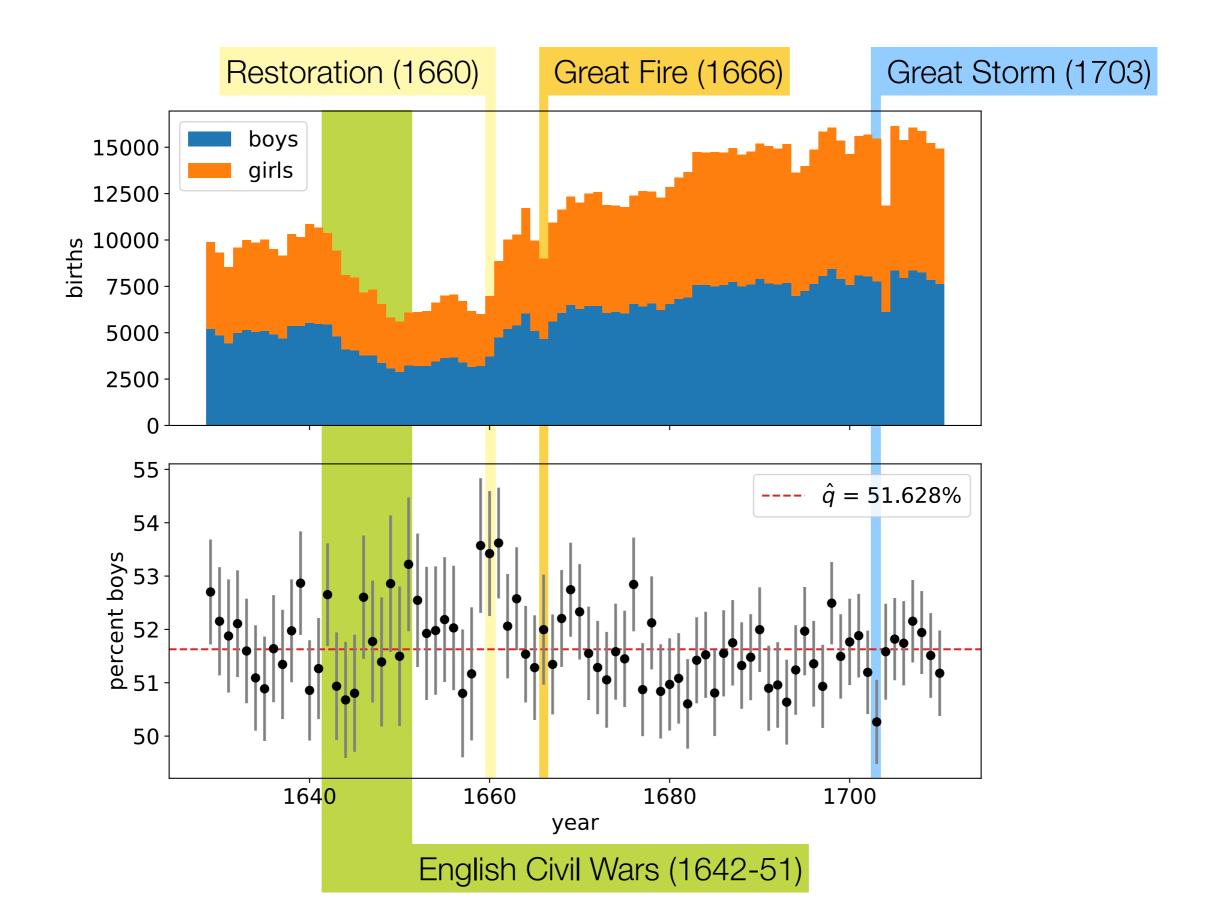
Row	Outcome	Observed %	Expected %
1	Round & yellow!	56.65	56.25
2	Round & green	19.42	18.75
3	Angular & yellow!	18.17	18.75
4	Angular & green!	5.76	6.25
Output			
Method to ca	alculate CI: Wilson/Brown (r	recommended)	\$
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+ New Graph	8	P value summary	ns			
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E O vs. E	11		Expected #	Observed #	Expected %	Observed %
	12	Round & yellow	312.8	315	56.25	56.65
	1:	Round & green	104.3	108	18.75	19.42
	14	Angular & yellow	104.3	101	18.75	18.17
	1	Angular & green	34.75	32	6.250	5.755
	16	TOTAL	556.0	556.0	100.0	100.00
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Example 4: Human sex ratio in London over time

Is it possible that the boy/girl ratio changes from year to year?



Chi square test (unknown proportions)

	male	female
1629	5218	4683
1630	4858	4457
1631	4422	4102
1632	4994	4590
1633	5158	4839
1634	5035	4820
1635	5106	4928
1636	4917	4605
1637	4703	4457
1638	5359	4952

sex

Null Hypothesis:

Two multi-category variables A and B are independent, i.e., $p(A, B) = p(A) \cdot p(B)$

Statistic:

$$\chi^2 = \sum \frac{(\text{observed} - \text{expected})^2}{\text{expected}}$$

Null distribution:Chi square distribution withDOF = nm - m - n + 1wherem = number of possible values for An = number of possible values or B

year

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Q~ Search		Table format:	Outcome A	Outcome B	Outcome C	Outcome D	Outcome E	Outcome F	Outcome G	Outcome I
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		17 1645	4047	3919						
		18 1646	3768	3395						
		19 1647	3796	3536						
		20 1648	3363	3181						
		21 1649	3079	2746						
		22 1650	2890	2722						
	-	23 1651	3231	2840						
		24 1652	3220	2908						
		25 1653	3196							
		26 1654	3441	3179						
		27 1655	3655	3349						
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		29 1657	3396	3289						
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▼ Transform, Normalize	A:boys
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Row many's with SD or SEM	
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 Parts of whole analyses 	
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Method to c	ompute the P	value			
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⊕ New Info	3	P value and statistical significance			
	> 4	Test	Chi-square		
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① New Analysis	6	P value	<0.0001		
	» 0 7		<0.0001 ****		
 ⊕ New Graph ▼ Layouts 		P value summary			
New Layout	0	One- or two-sided	NA		
	9	Statistically significant (P < 0.05)?	Yes		
	10				
	11	Data analyzed			
	12	Number of rows	82		
	13	Number of columns	2		
	14				
Family	> 15				
arbuthnot	16				
Contingency	17				
	18				
	19				
	20				
	21				
	22				
	22				
	24				
	25				
	26				
	27				
	28				
	29				
▲ ► €		Contingency of arbuthnot		• ص	